

JOINT SYNCHRONIZATION AND SYMBOL DETECTION IN ASYNCHRONOUS DS-CDMA SYSTEMS

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ABSTRACT

Design accurate estimators which also consider the noise term in low SNR scenarios is paramount to achieve optimal solutions and to obtain precise symbol detectors. Particularly, this paper estimates the propagation delays focusing on asynchronous DS - $CDMA$ systems. The proposed *Minimum Conditioned Variance (MCV)* is the choice in noisy environments, implementing the best linear detector of the transmitted symbols under a minimum mean-square error criterion. The result is an estimator that improves the *conditional ML (CML)* solution when noise is not negligible, and attains the derived *Gaussian Unconditional Cramér-Rao Bound (UCRB)* in the whole E_bN_0 range as classical *Gaussian Unconditional ML (UML)* does. Consequently, the proposed *MCV* estimator, becomes an optimal quadratic solution achieving similar features than *UML* in a straightforward way, and with no assumptions on the signal statistics.

1. INTRODUCTION

In digital communications, the knowledge of certain parameters as for example the phase and carrier frequency or the propagation delay, are paramount to get a reliable detection of the transmitted symbols. Focusing on multi-user DS - $CDMA$ systems, an accurate estimation of the propagation delays for all users is essential. Otherwise, the performance of the multi-user detector is rapidly decreased by means of multiple-access interference (MAI), as has been widely studied in the literature [1], [2]. Accordingly, this paper addresses a multi-parametric estimator intended for the multi-user synchronization and symbol detection, with high performance in low SNR scenarios. Nevertheless, the proposed algorithm is not restricted to multi-user synchronizers, and can be also extended to other estimation problems, like frequency synchronization in $OFDM$ and *Multi-Carrier* schemes.

Maximum Likelihood (ML) formulation has been usually employed to design timing estimators. Classically, *Unconditional ML (UML)* algorithms have been developed in

the field of digital communications modeling the transmitted symbols as stochastic processes. Nevertheless, in order to obtain feasible mathematical expressions, *UML* estimators make some assumptions on the gaussianity of signal statistics, which is known to be a non-realistic assumption in digital communications, or assumptions on low SNR , which leads to *self-noise* appreciable when the noise term is negligible. Consequently, the restrictions on *UML* motivated the introduction of *deterministic* or *conditional ML (CML)*, which considers the transmitted symbols as deterministic unknown parameters. This formulation has been applied by Stoica and Nehorai [3] in sensor array processing to perform DOA estimation, and more recently the same principle has been applied to frequency and timing estimation [4]-[7]. The *CML* solution does not present *self-noise*, is robust in *near-far* scenarios, and provides a high performance at high SNR 's. Nevertheless it is not an optimal solution in noisy scenarios with low SNR .

The proposed *Minimum Conditioned Variance (MCV)* method, addressed in this paper, mitigates the *CML* estimation drawbacks at low SNR scenarios considering the impact of the noise, and becomes the deterministic solution at high SNR . Although the derived *MCV* becomes biased, the bias value can be estimated and next subtracted to obtain an unbiased estimator. The result is an estimator that attains the lower *Gaussian Unconditional Cramér-Rao Bound UCRB* in the whole E_bN_0 range, as *Gaussian UML* does. Accordingly, *MCV* becomes an optimal quadratic estimator with no assumptions on the signal statistics.

This paper is organized as follows. Next section describes the discrete-time signal model, and obtains a structured matrix expression containing the parameters to estimate. Section 3 describes the *CML* formulation and justifies under which conditions the deterministic criterion does not become feasible. Afterwards, section 4 introduces the *Minimum Conditioned Variance* method as choice, and derives its gradient expression. Furthermore, a detailed study of the proposed estimator shows it is biased and consequently a modified unbiased estimator is proposed. Next, section 5 derives the *UCRB* which is used as a benchmark, at high and low SNR 's, to the performance of the proposed multi-user delay estimator. Finally last section presents some simulation results proving the proposed *MCV* outperforms *CML*, attaining the *UCRB* and reducing the Bit-Error Rate BER in symbol detection.

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2. DISCRETE-TIME SIGNAL MODEL

The described model considers a K user asynchronous DS-SS-CDMA system operating in a multipath environment. The received signal contains the superposition of K active users:

$$r(t) = \sum_{k=1}^K s^k(t - \tau_k) + w(t) \quad (1)$$

where $s^k(t)$ denotes the k -user received baseband signal, τ_k its the propagation delay, and $w(t)$ represents the received AWGN noise term with zero mean and variance σ_w^2 . For each user the received baseband signal is modeled as:

$$s^k(t) = \sum_{n=-\infty}^{\infty} d_n^k e^{j\theta_k} g^k(t - nT) \quad (2)$$

where $g^k(t)$ represents the k -user received signature, T is the bit duration, d_n^k are the transmitted information bits, and θ_k the received carrier phase. Moreover, considering the presence of a propagation multipath channel with baseband impulse response $h^k(t)$, the k -user received signature is given by a distorted version of the transmitted spreading waveform $c^k(t)$ as:

$$g^k(t) = c^k(t) * h^k(t) \quad (3)$$

Finally the received signal as a function of the user's signatures is given by:

$$r(t) = \sum_{k=1}^K \sum_{n=-\infty}^{\infty} d_n^k e^{j\theta_k} g^k(t - nT - \tau_k) + w(t) \quad (4)$$

The algorithm is derived in a discrete-time signal model by sampling the received waveform at N_{sc} samples per chip. Choosing the sampling frequency as $f_s = 1/T_s$, where T_s is the sampling period, and collecting $2M + 1$ samples of $r(nT_s)$, the vector \mathbf{r} can be defined as:

$$\mathbf{r} = [r(-MT_s) \quad \dots \quad r(0) \quad \dots \quad r(MT_s)]^T \quad (5)$$

At this point equation (4) can be expressed following the matrix signal model:

$$\mathbf{r} = \mathbf{A}(\tau)\mathbf{x} + \mathbf{w}^* \quad (6)$$

The set of unknown parameters (i.e. the transmitted symbols and phase errors) for k -user define the vector \mathbf{x}^k :

$$\mathbf{x}^k = [d_{-L}^k e^{j\theta_k} \quad \dots \quad d_0^k e^{j\theta_k} \quad \dots \quad d_L^k e^{j\theta_k}]^T \quad (7)$$

where the number of transmitted symbols $N_s = 2L + 1$. Finally, stacking all users, the nuisance parameter vector \mathbf{x} is defined as follows:

$$\mathbf{x} = [\mathbf{x}^{1T} \quad \mathbf{x}^{2T} \quad \dots \quad \mathbf{x}^{KT}]^T \quad (8)$$

^{*})The channel coefficients are assumed to be known or previously estimated (e.g. [8])

On the other hand the model transfer matrix, denoted as $\mathbf{A}(\tau)$ [†], contains the user signatures, and the parameters to estimate τ_k :

$$\mathbf{A} = \mathbf{A}(\tau) = [\mathbf{A}^1(\tau_1) \quad \mathbf{A}^2(\tau_2) \quad \dots \quad \mathbf{A}^K(\tau_K)] \quad (9)$$

$$\begin{aligned} \mathbf{A}^k(\tau_k) &= [\mathbf{a}_0^k \quad \mathbf{a}_1^k \quad \dots \quad \mathbf{a}_{N_s-1}^k] \\ \mathbf{a}_n^k &= [g^k(-MT_s - nT - \tau_k) \quad \dots \\ &\quad g^k(MT_s - nT - \tau_k)]^T \end{aligned}$$

where the columns of $\mathbf{A}^k(\tau_k)$ are scrolled versions of the k -user signature delayed τ_k .

A more detailed model of matrix $\mathbf{A}^k(\tau_k)$ will be constituted by the product of two matrices:

$$\mathbf{A}^k(\tau_k) = \mathbf{H}^k(\mathbf{h}_k) \mathbf{C}^k(\tau_k) \quad (10)$$

Matrix $\mathbf{H}^k(\mathbf{h}_k)$ is a Sylvester or convoluting matrix modeling the channel distortion, whose columns are the k -user impulsive channel response coefficients. On the other hand, matrix $\mathbf{C}^k(\tau_k)$ will be obtained by the k -user spreading code delayed τ_k .

3. THE CML FORMULATION

The cost function in CML estimation for the signal model in (6) is derived from the joint ML cost function that is formulated as:

$$\Lambda(\mathbf{r}/\tau, \mathbf{x}) = \frac{1}{(\pi\sigma_w^2)^M} e^{-\frac{1}{2\sigma_w^2} \|\mathbf{r} - \mathbf{A}\mathbf{x}\|^2} \quad (11)$$

The ML function depends on the parameter estimation vector τ and also on the vector \mathbf{x} . Notice that vector \mathbf{x} contains the set of unknown parameters and thus it is necessary to take some considerations on this vector. The joint τ, \mathbf{x} estimation could be the solution, but it is discarded because it is computationally complex, and alternative algorithms only focusing on the τ vector estimation are proposed. Classically, UML solution computes the expectation of the joint ML function with respect to the nuisance parameters:

$$\Lambda_{UML}(\mathbf{r}/\tau) = E_{\mathbf{x}} \{ \Lambda(\mathbf{r}/\tau, \mathbf{x}) \} \quad (12)$$

In general the expectation $E_{\mathbf{x}}$ in (12) is quite difficult to obtain, and in practice only an approximation of the likelihood function in low SNR scenarios is approached.

Previous limitations motivate the use of the CML solution. This method considers the nuisance parameters as deterministic, and thus they can be substituted by its estimation keeping fixed τ vector. The ML estimation of \mathbf{x} , when no restrictions are imposed on it, can be obtained as:

$$\hat{\mathbf{x}}_{ML} = \mathbf{A}^\# \mathbf{r} \quad (13)$$

where $\mathbf{A}^\#$ is the Moore-Penrose pseudo-inverse. Once the nuisance vector \mathbf{x} is estimated, the compressed ML function to maximize, which only depends on the parameter vector τ , is obtained by replacing (13) in (11). And finally the

[†])Hereafter the dependence on vector τ will be suppressed for simplicity

derived log-likelihood function to minimize (omitting irrelevant constants) is given by:

$$\min_{\tau} L_{CML}(\mathbf{r}/\tau) = \text{tr} \{ \mathbf{P}_A^\perp \hat{\mathbf{R}} \} \quad (14)$$

where $\mathbf{P}_A^\perp = \mathbf{I} - \mathbf{A}\mathbf{A}^\#$ is the projection matrix onto the orthogonal subspace defined by \mathbf{A} , and $\hat{\mathbf{R}} = \mathbf{r}\mathbf{r}^H$.

To minimize (14) a gradient algorithm may be used. The gradient in conditional *ML* was derived by Viberg, Ottersten and Kailath [9] in the context of array processing for *DOA* estimation. In our delay estimation problem this gradient can be expressed as:

$$g_{c_i} = -2\text{Re} \{ (\mathbf{r}^H \mathbf{P}_A^\perp \mathbf{D}_i) (\mathbf{A}^\# \mathbf{r}) \} \quad (15)$$

where $\mathbf{D}_i = \frac{\partial}{\partial \tau_i} \mathbf{A}$.

A more accurate study of the gradient expression shows that it is computed by the product of two terms. The first term is $(\mathbf{r}^H \mathbf{P}_A^\perp \mathbf{D}_i)$ and justifies the proposed algorithm to be *self-noise* free. Considering a noiseless environment, and the absence of delay errors, vector \mathbf{r} will be contained in the signal subspace generated by the \mathbf{A} matrix columns. Thus, the projection matrix \mathbf{P}_A^\perp , which does not appear in the classical unconditional approach, acts as a zero-forcer placed at the output of the derivative matched filter \mathbf{D}_i . As a result, the estimator ensures in all cases a *self-noise* free solution: $(\mathbf{r}^H \mathbf{P}_A^\perp \mathbf{D}_i) = 0$.

The second term $(\mathbf{A}^\# \mathbf{r})$ corresponds to the *ML* estimation of the unconstrained vector \mathbf{x} . Notice that this expression is the decorrelating detector solution, so the algorithm not only estimates the propagation delay but also implements this sub-optimum detector. The presence of this term justifies the proposed solution to be a robust *near-far* estimator. Analyzing the signal model (6) it is observed that the received powers can be introduced in the nuisance parameter vector \mathbf{x} . Hence, following (13) it is guaranteed that the algorithm will estimate the received power values, justifying the estimator to be insensitive to different power levels.

Nevertheless, the decorrelating detector evidences some difficulties in noisy scenarios. The pseudoinverse, as the ideal zero-forcing solution *ZF* in equalization, does not take into account the noise term. Accordingly, when the transfer matrix \mathbf{A} eigenvalue spreading, defined as:

$$\chi = \frac{\lambda_{A_{max}}}{\lambda_{A_{min}}} \quad (16)$$

is large enough, the noise term will be extremely increased, becoming the *CML* method an unacceptable solution in low *SNR* scenarios, which are common in wideband *DS-CDMA* systems.

4. MINIMUM CONDITIONED VARIANCE APPROACH

A novel approach is proposed in this paper considering the impact of the noise in the likelihood function, achieving in consequence a more robust estimator in low *SNR* scenarios. The *Minimum Conditioned Variance* approach (*MCV*) makes the nuisance parameter estimation as the best linear

estimation under a minimum variance criterion given an observation vector \mathbf{r} . This estimation is:

$$\begin{aligned} \hat{\mathbf{x}} &= E[\mathbf{x}/\mathbf{r}] = \mathbf{\Gamma} \mathbf{A}^H (\mathbf{A} \mathbf{\Gamma} \mathbf{A}^H + \sigma_w^2 \mathbf{I})^{-1} \mathbf{r} = \mathbf{C} \mathbf{r} \\ \mathbf{C} &= \mathbf{\Gamma} \mathbf{A}^H (\mathbf{A} \mathbf{\Gamma} \mathbf{A}^H + \sigma_w^2 \mathbf{I})^{-1} \\ \mathbf{\Gamma} &= E \{ \mathbf{x} \mathbf{x}^H \} \end{aligned} \quad (17)$$

Previous expression belongs to the best linear and non-linear estimator under Gaussian conditions, and only the best linear estimator under non-Gaussian conditions. The new cost function is derived by substituting (17) in equation (11) and it is given by:

$$\min_{\tau} L_{MCV}(\mathbf{r}/\tau) = \|\mathbf{r} - \mathbf{A} \mathbf{C} \mathbf{r}\|^2 \quad (18)$$

At high *SNR* scenarios $\mathbf{C}(\sigma_w^2 \rightarrow 0) = \mathbf{A}^\#$ is the pseudo-inverse of \mathbf{A} , becoming the *CML* solution. On the other hand, when the contribution of $\mathbf{A} \mathbf{\Gamma} \mathbf{A}^H$ is negligible in front of $\sigma_w^2 \mathbf{I}$, \mathbf{C} approaches a bank of matched filters containing all the user signatures: $\mathbf{C}(\sigma_w^2 \rightarrow \infty) = \sigma_w^{-2} \mathbf{\Gamma} \mathbf{A}^H$. This second limit is achieved at low *SNR* when the noise power is much greater than the received signal power for all users. Notice however that, in high *near-far* scenarios, the elements in $\mathbf{\Gamma}$ associated to the most powerful users will be higher than the noise term. Consequently, in scenarios with low *SNR* and small *near-far*, the *MCV* will improve the classical *CML* solution, whereas in high *near-far* scenarios, *MCV* will remain close to *CML*.

To minimize (18) we will follow once again a gradient scheme. The gradient expression in *MCV* is given by:

$$g_{m_{c_i}} = -2\text{Re} \left\{ \mathbf{r}^H (\mathbf{I} - \mathbf{A} \mathbf{C})^H \left(\mathbf{D}_i \mathbf{C} + \mathbf{A} \frac{\partial}{\partial \tau_i} \mathbf{C} \right) \mathbf{r} \right\} \quad (19)$$

It results interesting to analyze the behaviour at high and low *SNR* scenarios. At high *SNR* $\mathbf{C} \rightarrow \mathbf{A}^\#$, and making use of $\mathbf{P}_A^\perp \mathbf{A} = 0$, the second term in the previous gradient is asymptotically equal to zero: $\mathbf{r}^H (\mathbf{I} - \mathbf{A} \mathbf{C})^H \mathbf{A} \frac{\partial}{\partial \tau_i} \mathbf{C} \mathbf{r} = 0$. Thus the gradient becomes:

$$g_{m_{c_i}} (\sigma_w^2 \rightarrow 0) \simeq -2\text{Re} \left\{ \mathbf{r}^H (\mathbf{I} - \mathbf{A} \mathbf{C})^H (\mathbf{D}_i \mathbf{C}) \mathbf{r} \right\} \quad (20)$$

Likewise, at low *SNR*'s $\mathbf{C} \rightarrow \sigma_w^{-2} \mathbf{\Gamma} \mathbf{A}^H$, and the two components in the gradient (19) supply the same value. Hence, the asymptotic gradient derived in noisy environments corresponds with:

$$g_{m_{c_i}} (\sigma_w^2 \rightarrow \infty) \simeq -\frac{4}{\sigma_w^2} \text{Re} \left\{ \mathbf{r}^H \mathbf{D}_i \mathbf{\Gamma} \mathbf{A}^H \mathbf{r} \right\} \quad (21)$$

Notice that the second term can be dropped in both cases without losing information by the gradient.

Finally, for the special case when there is only one parameter to estimate, e.g. timing or frequency estimation in linear and non-linear modulations, another argument to eliminate the second term is detailed in [7] and next outlined. Considering that vector \mathbf{r} follows a Gaussian distribution, which is known to be a non-realistic assumption in digital communications, the Gaussian *UML* cost function becomes [†]:

$$\begin{aligned} L_{UML}(\mathbf{r}/\tau) &= \mathbf{r}^H \mathbf{R}^{-1} \mathbf{r} \\ \mathbf{R} &= (\mathbf{A} \mathbf{\Gamma} \mathbf{A}^H + \sigma_w^2 \mathbf{I}) \end{aligned} \quad (22)$$

[†]only applicable if $\mathbf{A}^H \mathbf{A}$ does not depend on the parameter to estimate

and the Gaussian *UML* gradient in previous equation is given by:

$$g_{uml_i} = -2Re \{ \mathbf{r}^H (\mathbf{I} - \mathbf{A}\mathbf{C})^H \mathbf{D}_i \mathbf{C} \mathbf{r} \} \quad (23)$$

Comparing last equation with (19), a further justification for removing the second term is obtained. Accordingly, in uni-parametric estimators, and assuming a Gaussian distribution for the transmitted symbols, the *MCV* gradient becomes the Gaussian *UML* gradient. Nevertheless, in multi-parametric estimators, the Gaussian *UML* cost function becomes more complex:

$$L_{UMLG}(\mathbf{r}/\tau) = \ln |\mathbf{R}| + \mathbf{r}^H \mathbf{R}^{-1} \mathbf{r} \quad (24)$$

Notice a new term $\ln |\mathbf{R}|$, which becomes constant in the uniparametric estimators when $\mathbf{A}^H \mathbf{A}$ does not depend on the parameter to estimate, is introduced. The gradient expression, derived in [3] cannot be identified with (19) anymore.

After the previous analysis, the *MCV* gradient can be asymptotically rewritten as:

$$g_{mcv_i} \approx -2Re \{ \mathbf{r}^H (\mathbf{I} - \mathbf{A}\mathbf{C})^H \mathbf{D}_i \mathbf{C} \mathbf{r} \} \quad (25)$$

A more accurate analysis of the previous gradient shows it is biased. It can be seen that in the absence of timing errors the gradient does not become the null vector. Therefore, the bias expression can be obtained computing the gradient expected value when the estimated timing vector equals the real timing vector:

$$\begin{aligned} Bias_i &= E \{ g_{mcv_i} \} |_{\hat{\tau}=\tau} \\ &= -2Re \{ Tr \{ \mathbf{\Gamma} \mathbf{A}_\tau^H (\mathbf{I} - \mathbf{A}_\tau \mathbf{C}_\tau)^H \mathbf{D}_{i_\tau} \} \} \end{aligned} \quad (26)$$

denoting \mathbf{A}_τ , \mathbf{C}_τ , \mathbf{D}_{i_τ} , the dependence of matrices on τ . Unfortunately previous expression cannot be computed by the estimator because the real timing vector τ is not a priori known. Nevertheless, the gradient expected value close to the real timing vector does not depend on the absolute timing error $\tau - \hat{\tau}$. Hence, an accurate bias estimation can be obtained if the estimated timing vector is used to compute (26):

$$\widehat{Bias}_i = -2Re \{ Tr \{ \mathbf{\Gamma} \mathbf{A}_{\hat{\tau}}^H (\mathbf{I} - \mathbf{A}_{\hat{\tau}} \mathbf{C}_{\hat{\tau}})^H \mathbf{D}_{i_{\hat{\tau}}} \} \} \quad (27)$$

As a result, an unbiased estimation of τ vector can be obtained according to a modified gradient where the bias is subtracted:

$$g_{mcv_i}^{unbiased} = -2Re \{ \mathbf{r}^H (\mathbf{I} - \mathbf{A}\mathbf{C})^H \mathbf{D}_i \mathbf{C} \mathbf{r} \} - \widehat{Bias}_i \quad (28)$$

5. PERFORMANCE ANALYSIS

This section derives the Gaussian Unconditional Cramér-Rao Bound (*UCRB*) to compare it with the proposed *CML* and *MCV* multi-user delay estimators analyzing its performance. As it is shown in [3] the *UCRB* is a valid lower bound for the variance of any consistent estimator based on the data sample covariance matrix.

As derived in [10] the ij th Fisher Information Matrix (*FIM*) element can be obtained as:

$$\{FIM_u\}_{ij} = Tr \{ \mathbf{R}^{-1} \mathbf{R}_i \mathbf{R}^{-1} \mathbf{R}_j \} \quad (29)$$

where:

$$\begin{aligned} \mathbf{R} &= E \{ \mathbf{r} \mathbf{r}^H \} = \mathbf{A} \mathbf{\Gamma} \mathbf{A}^H + \sigma_w^2 \mathbf{I} \\ \mathbf{R}_i &= \frac{\partial}{\partial \tau_i} \mathbf{R} \end{aligned} \quad (30)$$

Focusing on our estimation problem, assuming that the noise power is *a priori* known, (which is considered in the *MCV* case), and modeling the transmitted symbols to be zero mean independent random variables (i.e. $\mathbf{\Gamma}$ is a diagonal matrix) \mathbf{R}_i results:

$$\begin{aligned} \mathbf{R}_i &= \sigma_w^2 (\mathbf{D}_i \mathbf{A}^H + \mathbf{A} \mathbf{D}_i^H) \quad i = 1 \dots N_s \\ \mathbf{D}_i &= \frac{\partial}{\partial \tau_i} \mathbf{A} \end{aligned} \quad (31)$$

which can be substituted into (29) to obtain the *UCRB*.

6. SIMULATION RESULTS

To evaluate the *CML* (15) and *MCV* (28) estimators its performance was compared computing the Root-Mean Square Error (*RMSE*) in the timing delay estimation, and the Bit-error Rate (*BER*) in the symbol detection. Simulations were done considering 5 users, the spreading codes were Gold sequences with 7 chips per bit, the pulse shaping was a square-root raised cosine pulse with roll-off factor equal to 0.5 and the considered modulation was BPSK, and the over-sampling factor was $N_{sc} = 2$. Denoting B_L as the equivalent noise loop bandwidth, this parameter is related with the number of transmitted symbols as [12]:

$$B_L T = \frac{1}{2N_s} \quad (32)$$

and the *UCRB* lower bound is usually written as a function this bandwidth factor.

Figure 1 compares the proposed *MCV* versus the classical *CML* algorithm, and compares the *RMSE* with the derived *UCRB* lower bound assuming that the noise power σ_w^2 is *a priori* known (29) - (31). A low *SNR* scenario with *near-far* $NF=0$ and only one path per user on *AWGN* (i.e. no channel assumption) was simulated. As it can be seen in figure 1, due to the high eigenvalue spread, at low *SNR* the *CML* is not an optimal solution and does not achieve the derived *UCRB*. Under those conditions, the proposed *MCV* outperforms the *CML* algorithm and attains the *UCRB*, becoming a quadratic optimal solution. Figure 1 also illustrates how at high *SNR* the *MCV* becomes the *CML* solution, and asymptotically both attain the Cramér-Rao Bound.

A second simulation shows the performance of both algorithms in symbol detection, and illustrates once again the importance of *MCV* in noisy environments. Figure 2 compares the *BER* according to the ML estimation of vector \mathbf{x} (13) considered in *CML* estimation, and the MMSE estimator (17) introduced in the *MCV* approach. In order to illustrate the eigenvalues spread importance, two simulations, using 7 chips per bit spreading codes (associated eigenvalue spreading $\chi : 35$) and 15 chips per bit spreading codes (associated eigenvalue spreading $\chi : 6.25$), were done. As it can be seen, the higher the eigenvalue spreading is, the worse the *CML* solution performs. When the system

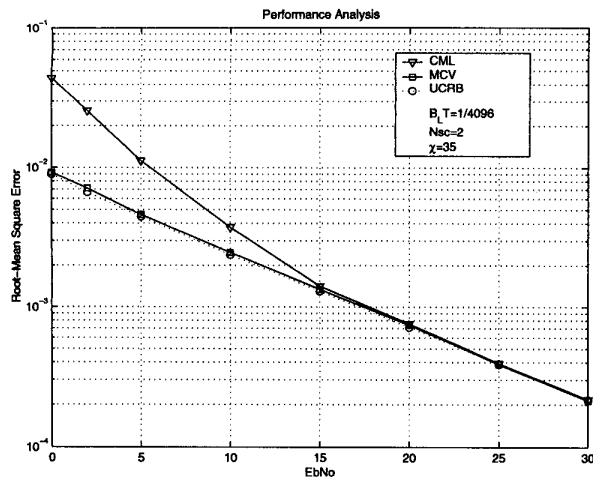


Figure 1: Timing Delay Estimation Error

is working at the limit of its capacity, (i.e. 5 users and spreading factor 7) the noise power is extremely increased by the decorrelating detector, and *CML* is not an acceptable solution, while the novel *MCV* always achieves a better performance.

7. CONCLUSIONS

In this paper the *MCV* algorithm has been introduced in the multiuser propagation delay estimation context. This novel method modifies the classical *CML* solution considering the impact of the noise in the Likelihood function compression. Hence, a more robust algorithm in noisy environments when the transference matrix eigenvalue dispersion is large, can be derived.

Simulations have shown *MCV* outperforms the classical deterministic algorithm in noisy conditions, and it corresponds asymptotically with the *CML* at high *SNR*'s. The mean squared timing error and the bit-error rate at the symbol detection have been used to evaluate this performance. Accordingly, the suggested quadratic estimation technique is shown to be optimal since it attains the *UCRB* lower bound in the whole *EbNo* range, becoming a great substitute not only to *CML*, but also to *UML* because it achieves similar features in a straightforward way.

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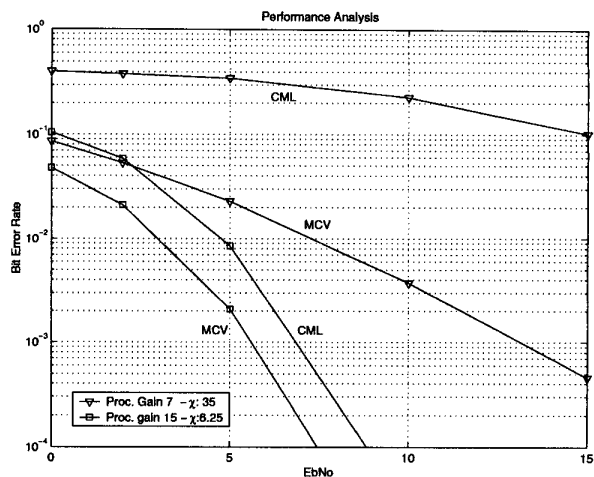


Figure 2: BER Symbol Detection

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